

TgY Calibration

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$$TgY(x_{TP}, y_{TP}, u_{TP}, q_{TP}) = \sum_{ijkl} a_{ijkl} x_{TP}^i y_{TP}^j u_{TP}^k q_{TP}^l =$$

Polynomial expansion.

$$= \left| \begin{array}{l} \text{We consider} \\ \text{25 matrix elements} \\ \text{for TgY} \end{array} \right| = a_{0000} \cdot 1 + a_{0001} \cdot q + a_{0003} \cdot q^3 +$$

$$+ a_{0010} \cdot y + a_{0011} \cdot q \cdot y + a_{0012} \cdot y \cdot q^2 + \dots$$

This is linear in a_{ijkl}

We search for matrix element using following eq.:

$$TgY^{(1)} \Big|_{HRSL}^{\text{event}} = a_{0000} \cdot 1 + a_{0001} \cdot q^{(1)} + a_{0003} \cdot q^{3(1)} + \dots$$

$$TgY^{(2)} \Big|_{HRSL} = a_{0000} \cdot 1 + a_{0001} \cdot q^{(2)} + a_{0003} \cdot q^{3(2)} + \dots$$

(Number of considered events)

This set of equations can be rewritten in matrix form as:

$$\begin{pmatrix} 1 & \varphi^{(1)} & \varphi^3(1) & Y^{(1)} & (\varphi Y)^{(1)} & \dots \\ 1 & \varphi^{(2)} & \varphi^3(2) & Y^{(2)} & (\varphi Y)^{(2)} & \dots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \end{pmatrix} \begin{pmatrix} a_{0000} \\ a_{0001} \\ a_{0003} \\ \vdots \\ \vdots \end{pmatrix} = \begin{pmatrix} Tg Y^{(1)} \\ Tg Y^{(2)} \\ \vdots \\ \vdots \end{pmatrix}$$



$$\boxed{10000 \times 25} \cdot \boxed{25} = \boxed{10000}$$

$$\underline{\underline{A}} \cdot \underline{Y} = \underline{b}$$

This is an over-determined system of equations. It can be solved using SVD.

Solution by Use of Singular Value Decomposition

In some applications, the normal equations are perfectly adequate for linear least-squares problems. However, in many cases the normal equations are very close to singular. A zero pivot element may be encountered during the solution of the linear equations (e.g., in `gaussj`), in which case you get no solution at all. Or a very small pivot may occur, in which case you typically get fitted parameters a_k with very large magnitudes that are delicately (and unstably) balanced to cancel out almost precisely when the fitted function is evaluated.

Why does this commonly occur? The reason is that, more often than experimenters would like to admit, data do not clearly distinguish between two or more of the basis functions provided. If two such functions, or two different combinations of functions, happen to fit the data about equally well — or equally badly — then the matrix $[A]$, unable to distinguish between them, neatly folds up its tent and becomes singular. There is a certain mathematical irony in the fact that least-squares problems are *both* overdetermined (number of data points greater than number of parameters) *and* underdetermined (ambiguous combinations of parameters exist); but that is how it frequently is. The ambiguities can be extremely hard to notice *a priori* in complicated problems.

Enter singular value decomposition (SVD). This would be a good time for you to review the material in §2.6, which we will not repeat here. In the case of an overdetermined system, **SVD produces a solution that is the best approximation in the least-squares sense, cf. equation (2.6.10).** That is exactly what we want. In the case of an underdetermined system, SVD produces a solution whose values (for us, the a_k 's) are smallest in the least-squares sense, cf. equation (2.6.8). That is also what we want: When some combination of basis functions is irrelevant to the fit, that combination will be driven down to a small, innocuous, value, rather than pushed up to delicately canceling infinities.

In terms of the design matrix A (equation 15.4.4) and the vector b (equation 15.4.5), minimization of χ^2 in (15.4.3) can be written as

$$\text{find } \mathbf{a} \text{ that minimizes } \chi^2 = |\mathbf{A} \cdot \mathbf{a} - \mathbf{b}|^2 \quad (15.4.16)$$

Comparing to equation (2.6.9), we see that this is precisely the problem that routines `svdcmp` and `svbksb` are designed to solve. The solution, which is given by equation (2.6.12), can be rewritten as follows: If U and V enter the SVD decomposition of A according to equation (2.6.1), as computed by `svdcmp`, then let the vectors $U_{(i)}$ $i = 1, \dots, M$ denote the *columns* of U (each one a vector of length N); and let the vectors $V_{(i)}$; $i = 1, \dots, M$ denote the *columns* of V (each one a vector of length M). Then the solution (2.6.12) of the least-squares problem (15.4.16) can be written as

$$\mathbf{a} = \sum_{i=1}^M \left(\frac{U_{(i)} \cdot \mathbf{b}}{w_i} \right) V_{(i)} \quad (15.4.17)$$

where the w_i are, as in §2.6, the singular values calculated by `svdcmp`.

Equation (15.4.17) says that the fitted parameters \mathbf{a} are linear combinations of the columns of V , with coefficients obtained by forming dot products of the columns

of \mathbf{U} with the weighted data vector (15.4.5). Though it is beyond our scope to prove here, it turns out that the standard (loosely, "probable") errors in the fitted parameters are also linear combinations of the columns of \mathbf{V} . In fact, equation (15.4.17) can be written in a form displaying these errors as

$$\mathbf{a} = \left[\sum_{i=1}^M \left(\frac{\mathbf{U}_{(i)} \cdot \mathbf{b}}{w_i} \right) \mathbf{V}_{(i)} \right] \pm \frac{1}{w_1} \mathbf{V}_{(1)} \pm \cdots \pm \frac{1}{w_M} \mathbf{V}_{(M)} \quad (15.4.18)$$

Here each \pm is followed by a standard deviation. The amazing fact is that, decomposed in this fashion, the standard deviations are all mutually independent (uncorrelated). Therefore they can be added together in root-mean-square fashion. What is going on is that the vectors $\mathbf{V}_{(i)}$ are the principal axes of the error ellipsoid of the fitted parameters \mathbf{a} (see §15.6).

It follows that the variance in the estimate of a parameter a_j is given by

$$\sigma^2(a_j) = \sum_{i=1}^M \frac{1}{w_i^2} [\mathbf{V}_{(i)}]_j^2 = \sum_{i=1}^M \left(\frac{V_{ji}}{w_i} \right)^2 \quad (15.4.19)$$

whose result should be identical with (15.4.14). As before, you should not be surprised at the formula for the covariances, here given without proof,

$$\text{Cov}(a_j, a_k) = \sum_{i=1}^M \left(\frac{V_{ji} V_{ki}}{w_i^2} \right) \quad (15.4.20)$$

We introduced this subsection by noting that the normal equations can fail by encountering a zero pivot. We have not yet, however, mentioned how SVD overcomes this problem. The answer is: If any singular value w_i is zero, its reciprocal in equation (15.4.18) should be set to zero, not infinity. (Compare the discussion preceding equation 2.6.7.) This corresponds to adding to the fitted parameters \mathbf{a} a *zero* multiple, rather than some random large multiple, of any linear combination of basis functions that are degenerate in the fit. It is a good thing to do!

Moreover, if a singular value w_i is nonzero but very small, you should also define *its* reciprocal to be zero, since its apparent value is probably an artifact of roundoff error, not a meaningful number. A plausible answer to the question "how small is small?" is to edit in this fashion all singular values whose ratio to the largest singular value is less than N times the machine precision ϵ . (You might argue for \sqrt{N} , or a constant, instead of N as the multiple; that starts getting into hardware-dependent questions.)

There is another reason for editing even *additional* singular values, ones large enough that roundoff error is not a question. Singular value decomposition allows you to identify linear combinations of variables that just happen not to contribute much to reducing the χ^2 of your data set. Editing these can sometimes reduce the probable error on your coefficients quite significantly, while increasing the minimum χ^2 only negligibly. We will learn more about identifying and treating such cases in §15.6. In the following routine, the point at which this kind of editing would occur is indicated.

Generally speaking, we recommend that you always use SVD techniques instead of using the normal equations. SVD's only significant disadvantage is that it requires

an extra array of size $N \times M$ to store the whole design matrix. This storage is overwritten by the matrix U . Storage is also required for the $M \times M$ matrix V , but this is instead of the same-sized matrix for the coefficients of the normal equations. SVD can be significantly slower than solving the normal equations; however, its great advantage, that it (theoretically) *cannot fail*, more than makes up for the speed disadvantage.

In the routine that follows, the matrices u, v and the vector w are input as working space. The logical dimensions of the problem are n data points by m basis functions (and fitted parameters). If you care only about the values a of the fitted parameters, then u, v, w contain no useful information on output. If you want probable errors for the fitted parameters, read on.

```
#include "nrutil.h"
#define TOL 1.0e-5
```

```
void svdfit(float x[], float y[], float sig[], int ndata, float a[], int ma,
           float **u, float **v, float w[], float *chisq,
           void (*funcs)(float, float [], int))
```

Given a set of data points $x[1..ndata], y[1..ndata]$ with individual standard deviations $sig[1..ndata]$, use χ^2 minimization to determine the coefficients $a[1..ma]$ of the fitting function $y = \sum_i a_i \times afunc_i(x)$. Here we solve the fitting equations using singular value decomposition of the n data by m matrix, as in §2.6. Arrays $u[1..ndata][1..ma]$; $v[1..ma][1..ma]$, and $w[1..ma]$ provide workspace on input; on output they define the singular value decomposition, and can be used to obtain the covariance matrix. The program returns values for the m fit parameters a , and χ^2 , $chisq$. The user supplies a routine $funcs(x, afunc, ma)$ that returns the m basis functions evaluated at $x = x$ in the array $afunc[1..ma]$.

```
{
    void svbksb(float **u, float w[], float **v, int m, int n, float b[],
               float x[]);
    void svdcmp(float **a, int m, int n, float w[], float **v);
    int j,i;
    float wmax,tmp,thresh,sum,*b,*afunc;

    b=vector(1,ndata);
    afunc=vector(1,ma);
    for (i=1;i<=ndata;i++) {
        (*funcs)(x[i],afunc,ma);
        tmp=1.0/sig[i];
        for (j=1;j<=ma;j++) u[i][j]=afunc[j]*tmp;
        b[i]=y[i]*tmp;
    }
    svdcmp(u,ndata,ma,w,v);
    wmax=0.0;
    for (j=1;j<=ma;j++)
        if (w[j] > wmax) wmax=w[j];
    thresh=TOL*wmax;
    for (j=1;j<=ma;j++)
        if (w[j] < thresh) w[j]=0.0;
    svbksb(u,w,v,ndata,ma,b,a);
    *chisq=0.0;
    for (i=1;i<=ndata;i++) {
        (*funcs)(x[i],afunc,ma);
        for (sum=0.0,j=1;j<=ma;j++) sum += a[j]*afunc[j];
        *chisq += (tmp=(y[i]-sum)/sig[i],tmp*tmp);
    }
    free_vector(afunc,1,ma);
    free_vector(b,1,ndata);
}
```

Accumulate coefficients of the fitting matrix.

Singular value decomposition.
Edit the singular values, given TOL from the #define statement, between here ...

...and here.

Evaluate chi-square.